

Pharmacogenetics of Tardive Dyskinesia: Combined Analysis of 780 Patients Supports Association with Dopamine D3 Receptor Gene Ser9Gly Polymorphism

Bernard Lerer, M.D., Ronnen H. Segman, M.D., Heiner Fangerau, M.D., Ann K. Daly, Ph.D., Vincenzo S. Basile, B.Sc., Roberto Cavallaro M.D., Harald N. Aschauer M.D., Robin G. McCreadie, M.D., Stephanie Ohlraun, M.Psych., Nicol Ferrier, M.D., Mario Masellis M.Sc., Massimiliano Verga, M.D., Joachim Scharfetter, M.D., Marcella Rietschel, M.D., Roger Lovlie, Ph.D., Uriel Heresco Levy, M.D., Herbert Y. Meltzer, M.D., James L. Kennedy, M.D., Vidar M. Steen, Ph.D., and Fabio Macciardi, M.D., Ph.D.

Variability among individuals in their therapeutic response to psychotropic drugs and in susceptibility to adverse effects is considerable. Pharmacogenetics addresses the contribution of genetic factors to this variability. An important focus of interest in pharmacogenetics has been on candidate genes that play a role in susceptibility to the antipsychotic drug-induced adverse effect, tardive dyskinesia (TD). Four published studies have reported an association between a serine (ser) to glycine (gly) polymorphism in exon 1 of the dopamine D3 receptor gene (DRD3) and TD; three failed to replicate this finding and one found an insignificant trend. We examined the association in a pooled sample of 780 patients (317 with TD and 463 without TD) drawn from 6 research centers, who were divided into 8 groups based on their population origin. The analysis employed stepwise logistic regression so as to

allow confounding effects of group, age, and gender to be taken into account. TD was significantly associated with DRD3 gly allele carrier status ($\chi^2=4.46$, df 1, p = .04) and with DRD3 genotype ($\chi^2=6.62$, df 2, p = .04) over and above the effect of group. Similar positive effects were observed when controlling for age and gender ($\chi^2=5.02$, df 1, p = .02 for gly allele carrier status; χ^2 = 7.51, df 2, p = .002 for genotype). Examining abnormal involuntary movement scores as a continuous variable, we found that patients homozygous for the gly allele had significantly higher scores than ser-gly heterozygotes (p = .006) or ser-ser homozygotes (p < .0001). We also performed a meta-analysis that included, besides the groups in the combined analysis, three other published studies on DRD3 and TD. The Mantel-Haenszel pooled odds ratio for DRD3 gly allele carrier status increasing susceptibility to TD was 1.33

From the Biological Psychiatry Laboratory, Department of Psychiatry, Hadassah-Hebrew University Medical Center, Jerusalem, Israel (BL, RHS); Department of Psychiatry, University of Bonn, Bonn, Germany (HF, SO, MR); Department of Pharmacological Sciences, University of Newcastle upon Tyne, Newcastle, UK (AKD); Department of Psychiatry, Centre for Addiction and Mental Health, Clarke Division, University of Toronto, Toronto, Canada (VSB, MM, JLK, FM); Department of Psychiatry, IRCCS H San Raffaele, Università Vita e Salute, Milan, Italy (RC, MV); Department of General Psychiatry, University Hospital for Psychiatry, Vienna, Austria (HNA, JS); Department of Clinical Research, Crichton Royal Hospital, Dumfries, UK (RGM); Department of Psychiatry, University of Newcastle upon Tyne, Newcastle, UK (NF); Dr. E. Martens Research Group for Biolog-

ical Psychiatry and Locus on Neuroscience, Center for Medical Genetics and Molecular Medicine, Haukeland University Hospital, University of Bergen, Norway (RL, VMS); Sara Herzog Memorial Hospital, Jerusalem, Israel (UHL); Vanderbilt University Medical Center, Psychiatric Hospital at Vanderbilt, Nashville, TN, USA (HYM).

Address correspondence to: Prof. Bernard Lerer, Director, Biological Psychiatry Laboratory, Department of Psychiatry Hadassah-Hebrew University Medical Center, Ein Karem, Jerusalem 91120, Israel. E-mail: lerer@cc.huji.ac.il

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(95% CI 1.04–1.70, $p = .02$); the cumulative pooled estimate showed an odds ratio of 1.52 (95% CI 1.08–1.68, $p < .0001$). These findings support a small but significant contribution of the DRD3 ser9gly polymorphism to TD susceptibility that is demonstrable over and above

population effects and the effect of age and gender on the phenotype.

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Currently, there is considerable interest in the application of emerging genomic technologies to the clinical practice of psychopharmacology (Catalano 1999; Veenstra-VanderWeele et al. 2000; Cichon et al. 2000; Pickar and Rubinow 2001; Lerer, in press). Variability among individuals in their degree of response to psychotropic drugs and in susceptibility to adverse effects is considerable. It is widely assumed that a substantial portion of this variability is genetically determined. Based on this assumption, identification of the genes involved and of pharmacogenetically relevant variation in their structure is regarded as an important research priority. The emphasis has been on functionally significant genetic variants, most often single nucleotide polymorphisms (SNPs) located in coding or regulatory regions of candidate genes that have a priori relevance to the drug response or adverse effect phenotype under study. SNPs located in intronic regions, however, can also have an impact on the coded protein (Krawezak et al. 1992). Although genetic factors that influence drug metabolism were studied in psychopharmacology long before current genomic technology was available (Poolsup et al. 2000; Steimer et al. 2001), work on pharmacodynamic aspects is still at an early stage. It has focused primarily on response to specific serotonin reuptake inhibitors (SSRIs) in major depression (Smeraldi et al. 1998, Lerer and Macciardi in press) and to atypical antipsychotic drugs, primarily clozapine, in schizophrenia (Arranz et al. 1998, Scharfetter et al. 1999a, Masellis et al. 2000).

An important focus of interest in pharmacogenetics has been on candidate genes that play a role in susceptibility to tardive dyskinesia (TD). TD is a choreoathetotic movement disorder with a chronic fluctuating course that affects 20% or more of chronic schizophrenia patients who are treated with classical, dopamine D2 receptor antagonist, antipsychotic drugs (Kane and Smith 1982; Jeste and Wyatt 1982; Kane 1995). Age is the strongest-known risk factor for the development of TD (Smith and Baldessarini 1980; Jeste et al. 1995). Other factors include duration and intensity of prior antipsychotic exposure, female gender, organic brain abnormalities, smoking, and affective disorder (Kane and Smith, 1982; Kane et al. 1992; Yassa and Jeste 1992). As demographic and drug-related variables predict only a minor part of the variance in the incidence of TD, a

prominent pharmacogenetic component appears plausible. Sporadic clinical reports of aggregation of TD cases in families (Yassa and Ananth 1981; Weinhold et al. 1981; Waddington and Yousseff 1988; Yousseff et al. 1989) and strain differences in the susceptibility of rodents to antipsychotic-induced repetitive jaw movements and vacuous chewing (Rosengarten et al. 1994; Tamminga et al. 1990), provide some support for a genetic predisposition. There are no studies that have explored the possible mode of inheritance of predisposition to TD as a genetic trait. Such studies are very difficult to perform because level of exposure to antipsychotic drugs and the precise drug administered to individual family members cannot be controlled. For the same reason, linkage analyses are not an applicable approach in the search for predisposing genes. Studies of the molecular genetics of TD have thus far employed a case-control association design implemented in samples of unrelated patients.

Attempts to elucidate pharmacogenetic factors in TD are important at two levels. First, in spite of the increasing use of atypical antipsychotic drugs, which have a lesser propensity to induce TD, classical antipsychotic agents are still widely used worldwide (Emsley et al. 1999). Second, TD is a relatively well-defined, semi-quantitative phenotype. As a model for pharmacogenetic studies in general, analysis of the molecular genetic basis of individual susceptibility to TD can provide important insights into crucial methodologic and statistical issues that are of general relevance to pharmacogenetics.

A number of genes have been studied as potential contributors to individual sensitivity to TD. These include genes implicated in the metabolism of antipsychotic drugs such as the genes for cytochrome P-450 2D6 (Arthur et al. 1995; Armstrong et al. 1997; Andreasen et al. 1997, Kapitany et al. 1998; Ohmori et al. 1999) and cytochrome P-450 1A2 (Basile et al. 2000), genes coding for the dopamine D2 (Chen et al. 1997) and D3 receptors (see below); serotonergic receptor genes such as the 5-HT_{2C} receptor (Segman et al. 2000) and the 5-HT_{2A} receptor (Segman et al. 2001; Basile et al. 2001, Tan et al. 2001), the serotonin transporter gene (Chong et al. 2000), and the manganese superoxide dismutase (MnSOD) gene (Hori et al. 2000). Earlier studies with Human Leukocyte Antigen (HLA) subtypes should also be noted (Canoso et al. 1986; Meltzer et al. 1990; Brown and White, 1991). (See Segman and Lerer, in press, for a review of candidate gene studies in the pharmacogenetics of TD).

The largest series of studies on candidate genes for TD has focused on the dopamine D3 receptor (DRD3). The human DRD3 gene has been localized to chromosome 3q13.3 by in situ hybridization (Le Coniat et al. 1991). D3 receptors readily bind classic and also atypical antipsychotic drugs but differ from other dopamine receptor subtypes in that they are primarily localized to limbic regions that may be particularly important in the regulation of emotions and in the pathogenesis of schizophrenia (Sokoloff et al. 1990; Zahm and Brog 1992). DRD3 contains a polymorphic site in the first exon that gives rise to a serine (ser) to glycine (gly) substitution in the N-terminal extracellular domain of the receptor protein (Lannfelt et al. 1992). The in vivo functional significance of the polymorphic ser9gly site is unknown. DRD3 receptor-binding analysis of Chinese hamster ovary (CHO) cells infected with Semliki Forest Virus to express either the wild-type cDNA or a recombinant ser9gly or both showed similar pharmacologic properties for several D3 receptor ligands. The cells expressing the DRD3gly variant, however, showed a sig-

nificantly higher binding affinity only for dopamine, whereas heterozygotes (i.e., doubly infected cells) were not significantly different from the wild type. In addition, both DRD3gly-gly homozygotes and DRD3ser-gly heterozygotes showed significantly higher binding affinity for the selective D3 ligand GR99841, compared with the wild-type receptor (Lundstrom & Turpin 1996). Although these results do not allow a straightforward extrapolation with regard to the biological significance of heterozygote versus homozygote status, they speak for either examining DRD3gly-gly homozygotes against all other genotypes or for grouping DRD3gly-gly homozygotes and DRD3ser-gly heterozygotes against wild-type homozygotes. Such extrapolations should be taken with caution, however, as these results were obtained in an in vitro recombinant system not necessarily reflecting in vivo receptor status in the brain. A clearer understanding of the functional significance of the ser9gly site must await more detailed studies employing in vivo radioligand binding, as well as postmortem brain and in vitro receptor studies.

Table 1. Published studies on the DRD3 ser9gly polymorphism and TD

Publication	Subjects	Findings from categorical analysis	Findings from continuous analysis
Steen et al. (1997)	TD-Y: 51 TD-N: 49 Origin: Caucasian	Excess of gly allele ($p = .035$) in TD Excess of gly-gly genotypes ($p = .018$) in TD	Not reported
Inada et al. (1997)	TD-Y: 49 TD-N: 56 Origin: Japanese	No significant difference in allele or genotype frequency	Not reported
Segman et al. (1999)	TD-Y: 53 TD-N: 63 Control: 117 Origin: Israeli Ashkenazi and non-Ashkenazi Jews	Excess of gly allele ($p = .055$) in TD Excess of ser-gly genotypes ($p = .0008$) in TD	Higher AIMS total ($p = .02$), orofacial ($p = .052$), trunk ($p = .p = .01$), and incapacitation ($p = .03$) scores in patients carrying the gly allele (ser-gly and gly-gly genotypes vs. ser-ser genotypes)
Basile et al. (1999)	TD categorization not reported. Origin: Caucasian, $n = 85$, African American, $n = 25$, Asian, $n = 2$	Not reported	Higher AIMS total scores ($p = .0005$) in patients with gly-gly genotypes vs. ser-ser genotypes and ser-gly genotypes. (Also in Caucasians and African Americans analyzed separately)
Rietschel et al. (2000)	TD-Y: 79 TD-N: 78 Origin: German	No significant difference in allele or genotype frequency	Not reported
Lovlie et al. (2000)	TD-Y: 32 TD-N: 39 Origin: Caucasian	No significant difference in allele or genotype frequency	Not reported
Garcia-Barcelo et al. (2001)	TD-Y: 65 TD-N: 66 Origin: Chinese (Hong Kong)	No significant difference in allele or genotype frequency	Not reported
Liao et al. 2001	TD-Y: 21 TD-N: 94 Origin: Chinese (Taiwan)	Allele frequencies not compared. Excess of ser-gly genotypes ($p = .009$) in TD.	Higher AIMS total scores in patients with ser-gly genotypes compared to patients with ser-ser and gly-gly genotypes ($p = .014$)

Tardive dyskinesia (TD).

Abnormal Involuntary Movement Scale (AIMS).

A summary of published studies of the DRD3 gene as a risk factor for TD is given in Table 1. Steen et al. (1997) first reported association of the DRD3 gly allele and TD in schizophrenia patients and observed a significant excess of DRD3gly-gly homozygotes among the patients with TD. The finding of Steen et al. (1997) has subsequently been supported by the observations of Segman et al. (1999), Basile et al. (1999), Liao et al. (2001), and (at a trend level) by Lovlie et al. (2000), but not by Inada et al. (1997), Rietschel et al. (2000), and Garcia-Barcelo et al. (2001). Important differences among these findings should be noted. Whereas the findings of Steen et al. (1997) suggested homozygosity for the DRD3 gly allele as a risk factor for TD, the findings of Segman et al. (1999) pointed to a dominant model and demonstrated an excess of DRD3gly carriers, whether as homo- or heterozygotes, among patients with TD. The findings of Liao et al. (2001) were similar to those of Segman et al. (1999). Basile et al. (1999) did not report a categorical analysis of their data but presented their findings as significantly higher scores on the Abnormal Involuntary Movements Scale (AIMS) (Guy, 1976) in patients homozygous for the DRD3gly allele. Segman et al. (1999) presented their results parametrically as well as categorically, demonstrating significantly higher AIMS scores in patients carrying the DRD3 gly allele as homo- or heterozygotes versus patients homozygous for the wild type ser allele. Liao et al. (2001) also presented a parametric analysis, finding significantly higher AIMS scores in patients carrying the ser-gly genotype compared with carriers of ser-ser or gly-gly. A further point of note is that only Steen et al. (1997), Inada et al. (1997), and Rietschel et al. (2000) based their definition of TD on repeated assessments of abnormal involuntary movements in their patients, thus fulfilling the criteria of Schooler and Kane (1982) for definite TD in regard to the patients categorized as fulfilling the diagnosis.

While suggesting a possible role for DRD3 in conferring genetic susceptibility to TD, the literature leaves a number of crucial questions unanswered or answered only in part. Central among these is whether the observed effects can be generalized among different populations, and whether population differences contribute to the different results obtained. Further issues are the extent to which known risk factors for TD such as age and gender impact on the findings. Noteworthy in this regard, Segman and Lerer (2002) recently found that age might not only influence vulnerability to TD but the contribution of certain genetic polymorphisms to this vulnerability.

In the present article, we present the results of a combined analysis of original data from several published papers on DRD3 and TD. To these we added data from two samples that have been presented in abstract form but not published in full. Our analysis took into account

the possible confounding effects of the origin of the patients, differences in the rating of abnormal involuntary movements between centers, and the impact of age and gender in the context of a combined sample that had significantly greater statistical power than the component samples to address the role of DRD3 as a risk factor for TD.

MATERIALS AND METHODS

Patients

Demographic, clinical, and genotypic data on the patients included in this analysis were supplied by six research centers: the University of Bergen, Norway; University of Bonn, Germany; Hadassah-Hebrew University Medical Center, Jerusalem, Israel; San Raffaele Institute, Milan, Italy; University of Toronto, Canada; and University of Vienna, Austria. The minimum criteria for inclusion in the analysis were information on age and gender; diagnosis of schizophrenia or schizoaffective disorder according to DSM-IV or ICD-9; presence or absence of TD established according to the Research Diagnostic Criteria for TD (Schooler and Kane 1982) and genotypic status on the DRD3 ser9gly polymorphism. AIMS scores were available for patients supplied by the Bergen (Newcastle and Nithsdale samples), Bonn, Jerusalem, and Toronto centers. Patients from the Milan center were assessed for abnormal involuntary movement by the Rockland Simpson Scale (Cavallaro et al. 1993), and patients from Vienna by the Tardive Dyskinesia Rating Scale (Simpson et al. 1979). Twenty three patients from Vienna were not assessed for TD with a rating scale; the diagnosis was made on the basis of clinical examination. Key analyses were repeated without these patients, and results did not differ substantively from those reported here (available on request). Because TD ratings were performed twice in only some of the centers, categorization according to presence or absence of TD in the present analysis is based on a single evaluation (the first if there was more than one evaluation). The TD diagnosis is thus at the probable level according to the RDC-TD criteria (Schooler and Kane 1982).

For the current analysis, patients were divided into eight groups regarding which further details are provided below and in Table 2. Data on the Bonn (Rietschel et al. 2000), Jerusalem (Segman et al. 1999), Newcastle (Lovlie et al. 2000), Nithsdale (Steen et al. 1997), and Toronto African American and Toronto Caucasian groups (Basile et al. 1999) had been published previously. Data on the Milan (Macciardi et al. 1996, 1997) and on part of the patients of the Vienna group, included in a larger data set (Scharfetter et al. 1999b), had been reported in abstract form. There were 240 patients from Bonn in the current analysis as compared with 157

Table 2. Demographic and clinical details of the sample

Group	TD-Y				TD-N				Origin	TD assessment
	N	Age	Gender	Diagnosis	N	Age	Gender	Diagnosis		
Bonn	85	44.6 ± 11.6	F-41 M-44	SCZ-20 SA-65	155	39.4 ± 10.2	F-86 M-69	SCZ-42 SA-113	German	AIMS
Jerusalem	55	50.8 ± 10.3	F-26 M-29	SCZ-55	58	50.8 ± 10.3	F-28 M-30	SCZ-58	Israeli Ashkenazi (n = 66) and non-Ashkenazi (n = 47) Jews	AIMS
Milan	34	55.8 ± 10.4	F-16 M-18	SCZ-34	59	53.4 ± 10.6	F-30 M-29	SCZ-59	Northern Italian	Rockland Simpson Dyskinesia Scale
Newcastle [†]	32	53.1 ± 18.0	F-4 M-27	SCZ-32	37	41.6 ± 12.3	F-13 M-24	SCZ-37	Caucasian	AIMS
Nithsdale	51	56.2 ± 15.8	F-23 M-28	SCZ-51	49	47.9 ± 13.6	F-21 M-28	SCZ-49	Caucasian	AIMS
Toronto: African American	10	35.5 ± 11.5	F-3 M-7	SCZ-10	13	30.1 ± 11.3	F-5 M-8	SCZ-13	African American	AIMS*
Toronto: Caucasian	30	34.9 ± 8.3	F-11 M-19	SCZ-30	51	31.1 ± 8.8	F-11 M-40	SCZ-51	Caucasian	AIMS*
Vienna	20	37.7 ± 13.2		SCZ-11 SA-5	41	30.8 ± 10.0	F-18 M-23	SCZ-30 SA-15	Austrian	TDRS (n = 38) Clinical (n = 23)
			F-9 M-11							

female (F); male (M); schizoaffective (SA); schizophrenia (SCZ); Tardive Dyskinesia Rating Scale (TDRS).

* Abnormal Involuntary Movement Scale. Converted from modified Hillside Simpson Dyskinesia Scale

[†]Gender of 1 tardive dyskinesia (TD) subject not known

in the original report (Rietschel et al. 2000) because patients with a single evaluation for TD were included here. For Jerusalem, the number of subjects is 113 because 3 subjects were blindly excluded to achieve closer age matching. The Newcastle group consisted of patients recruited in Newcastle, United Kingdom, who were genotyped by the Bergen group (Lovlie et al. 2000). The Nithsdale group consisted of patients recruited in Nithsdale, Scotland, who were also genotyped in Bergen (Steen et al. 1997). The patients from the Toronto center had been recruited in three research clinics in the United States and were genotyped in Toronto. For the current analysis, Caucasian and African American patients were included as separate groups, and the small number of Asian American patients (n = 2) was omitted from the analysis. Also, subjects in the Toronto sample who did not have a categorical diagnosis of TD but only AIMS scores were excluded from the current analysis. Table 2 provides a breakdown of the groups by age and gender, according to presence or absence of a categorical diagnosis of TD.

Genotyping

The ser9gly polymorphism in DRD3 was genotyped in the laboratories of the six participating centers by PCR,

MSC I restriction endonuclease, and electrophoresis, according to the method of Lannfelt et al. (1992). The detailed procedures followed in the Bonn (Rietschel et al. 2000), Bergen (Steen et al. 1997; Lovlie et al. 2000), and Jerusalem centers (Segman et al. 1999) are described in their publications. The same procedures were followed in the Milan and Vienna laboratories.

Statistical Analysis

For categorical analyses, maximum likelihood chi square statistics were employed. For bivariate comparisons of continuous data, the student *t*-test was used, and for multivariate comparisons, analysis of variance (ANOVA) or covariance (ANCOVA). These analyses were performed with Statistica for Windows 4.5 (Statsoft Inc., 1993) or SPSS 10.0 for Windows (SPSS Inc., 2000). To evaluate the relationship of more than one independent variable to a single binomial-dependent variable, stepwise logistic regression procedures were implemented, as previously described by Lerer et al. (2001). The STATA 6.0 program (Stata Technical Bulletin # 48 [STB-48] 1999), a general statistical package, was employed for these analyses and also for the meta-analysis described below. Multinomial logistic regression (mlogit) was applied to analyze genotypes, considered as multilevel categorical variables.

The aim of our analyses was to reevaluate the potential association of TD with the DRD3 gene when the raw data were available through the collaboration of the centers that presented the original reports. The statistical analyses were designed to examine the cumulative predictability of DRD3 as a risk factor for TD while controlling for sources of heterogeneity that could explain conflicting results obtained in the various reports.

First, we performed a meta-analysis estimating and plotting the odds ratios of the contributing centers for TD given the DRD3 risk factor, thus obtaining a statistical estimate of the heterogeneity across centers (groups). Published reports that were not part of the current pooled analysis were included in this meta-analysis. Then, we implemented two sets of analyses on the pooled sample, to possibly explain the degree of heterogeneity. For the first set of analyses, the outcome (response) variables of the logistic regression were DRD3 gly allele carrier status or DRD3 ser-gly genotype. The covariates of the model were the clinical phenotype of interest (i.e., TD present versus absent and the group, given a possible population effect on the frequency of the DRD3 gly allele). In this case, the phenotype of interest is a predictor of outcome and group, a potential confounder. Group was tested by generating a set of dummy variables equal to the number of groups included in the sample in order to evaluate the specific effect of any group. Results from the model are presented as likelihood ratio tests. In a further set of analyses, the outcome variable of the logistic regression was TD phenotype (present or absent), and the covariates of the model were age, gender, and DRD3 gly allele carrier status or DRD3 ser-gly genotype distribution. *P* values <.05 (two tailed) were regarded as significant in all

analyses; *p* values <0.1 (two tailed) are given in order to indicate trends, and ≥ 0.1 are listed as NS.

RESULTS

Initially, we checked for Hardy Weinberg equilibrium (HWE) in each of the index (TD) and comparison (non-TD) groups. All the non-TD groups were in HWE. Among the TD groups, those from Jerusalem ($\chi^2 = 11.98$, $p = .0005$) and Nithsdale ($\chi^2 = 4.42$, $p = .04$) were not in HWE. We also compared DRD3 gly allele frequency across all the groups. There was a significant difference in allele frequency ($\chi^2 = 46.4$, $df\ 7$, $p < .0001$). This was due primarily to the Toronto African American group which had a DRD3 gly allele frequency of 0.80 as compared considerably lower frequencies in the other groups (Bonn: 0.32, Jerusalem: 0.35, Milan: 0.39, Newcastle: 0.38, Nithsdale: 0.31, Toronto Caucasians: 0.35, Vienna: 0.34). Excluding the Toronto African Americans, the difference in DRD3 ser9gly allele frequency among the groups was not significant ($\chi^2 = 4.7$, $df\ 6$, $p > .1$).

Table 3 shows a comparison of the frequency of DRD3 alleles between patients with and without TD in each of the groups. There was significant excess of DRD3 gly alleles among the TD patients in the Nithsdale ($p = .02$) and Toronto African American ($p = .02$) groups and a trend in this direction in the Jerusalem group ($p = .09$). In the Vienna group, the allelic effect was opposite in direction, with a significant excess of ser alleles in patients with TD ($p = .02$).

Table 3 also shows DRD3 ser-gly genotypes in the TD and non-TD patients of each group. Comparison of

Table 3. Allele and genotype frequencies in the individual groups

Group	TD Status	No. Patients	Allele Frequency		Significance		Genotype Frequency		Significance
			ser	gly	(ML χ^2 , $df\ 1$)	ser-ser	ser-gly	gly-gly	(LR χ^2 , $df\ 1$)
Bonn	TD-Y	85	0.66	0.34	NS	0.44	0.46	0.11	NS
	TD-N	155	0.69	0.31		0.47	0.44	0.90	
Jerusalem	TD-Y	55	0.59	0.41	$\chi^2 = 2.8$ $p = .09$	0.24	0.71	0.55	$\chi^2 = 15.9$ $p = .0004$
	TD-N	58	0.70	0.30		0.48	0.43	0.86	
Milan	TD-Y	34	0.60	0.40	NS	0.38	0.44	0.18	NS
	TD-N	59	0.62	0.38		0.34	0.56	0.10	
Newcastle	TD-Y	32	0.56	0.44	NS	0.34	0.44	0.22	NS
	TD-N	37	0.68	0.32		0.46	0.43	0.11	
Nithsdale	TD-Y	51	0.62	0.38	$\chi^2 = 5.3$ $p = .02$	0.45	0.33	0.22	$\chi^2 = 9.5$ $p = .009$
	TD-N	49	0.77	0.23		0.57	0.39	0.04	
Toronto: African American	TD-Y	10	0.50	0.95	$\chi^2 = 5.4$ $p = .02$	—	0.10	0.90	—*
	TD-N	13	0.31	0.69		—	0.62	0.38	
Toronto: Caucasian	TD-Y	30	0.63	0.37	NS	0.43	0.40	0.17	NS
	TD-N	51	0.66	0.34		0.37	0.57	0.06	
Vienna	TD-Y	20	0.80	0.20	$\chi^2 = 5.2$ $p = .02$	0.70	0.20	0.10	$\chi^2 = 7.6$ $p = .02$
	TD-N	41	0.60	0.40		0.41	0.37	0.22	

*Significance could not be computed because of absence of ser-ser homozygotes.

NS = not significant ($p \geq 0.1$).

genotype frequency was by the subroutine "genhwcci" as implemented in STATA 6.0 (Stata Technical Bulletin # 48 [STB-48] 1999). The program estimates allele frequency, genotype frequencies, and disequilibrium coefficients for co-dominant traits or data of completely known genotypes in case-control studies. It performs asymptotic Hardy-Weinberg equilibrium test for both cases and controls in case-control studies. It also tests the Hardy-Weinberg equilibrium for genotypic counts of cases, under the assumption that the genotypic counts of controls are under Hardy-Weinberg equilibrium. This test showed a significant difference in the Jerusalem group ($\chi^2 = 15.85$, $p = .0004$), caused by an excess of ser-gly heterozygotes among the TD patients. The genotypes of the Nithsdale TD patients were significantly different from those of the non-TD patients in this group ($p = .009$) because of an excess of gly-gly homozygotes. As there were no ser-ser homozygotes among the Toronto African American patients, the test could not be performed for this group. A simple chi square test, however, showed a significant excess of gly-gly homozygotes among the African American TD patients ($p = .008$). In the Vienna group, genotype frequencies were also significantly different between TD and non-TD patients ($\chi^2 = 7.61$, $df=2$, $p = .02$), but in this case, there was an excess of ser-ser homozygotes among the TD patients.

Figure 1 shows the results of the meta-analysis and presents observed distribution of odds ratios for TD given DRD3 gly presence in seven of the groups included in the pooled analysis (Toronto African Americans excluded), as well as in three additional published reports. The Mantel-Haenszel (M-H) pooled odds ratio across all the available data is 1.33 (95% CI 1.04–1.70), pointing to a small but significant effect of DRD3gly as risk factor for TD. The cumulative pooled estimate (Lau et al. 1992) shows a cumulative odds ratio of 1.52 (95% CI 1.08–1.68), corresponding to a p -value $<.0001$, suggesting a much stronger effect of DRD3 gly as a risk factor. It is important, however, to note the degree of heterogeneity ($\chi^2 = 16.97$, $df = 9$, $p = .05$), with odds ratios ranging from 0.3 for the Vienna groups to 3.53 for Liao et al. (2001). Therefore, a major goal was to disentangle the origin of the heterogeneity and to re-evaluate TD/DRD3gly association while controlling for potential confounders. These are: 1) differences among groups in DRD3 ser-gly allele frequency; 2) possible group differences in the evaluation of TD; and 3) the impact of age and gender, both known to be susceptibility factors for TD. Therefore, we implemented a series of stepwise logistic regression procedures.

In the first set of analyses, the response variable was DRD3 gly allele carrier status (ser-gly heterozygotes and gly-gly homozygotes defined as gly allele carriers). With stepwise logistic regression analysis, we evaluated the effect of group and TD status as predictors of

the response variable. The likelihood ratio (LR) test of the two models (group + TD status and group alone) determines the significance of TD status controlled for the potential confounding effect of group. The results of this analysis are shown in Table 4A. The LR test for TD status predicting DRD3 gly allele carrier status was significant ($\chi^2=4.46$, $df 1$, $p = .04$). The combined analysis of genotypes was conducted according to the same approach. A multinomial logistic regression with the dependent variable being DRD3 genotype (3 levels) showed a significant LR test for TD status ($\chi^2= 6.62$, $df 2$, $p = .04$), while controlling for the possible confounding effects of populations, as shown in Table 4B.

In the second set of logistic regressions, the response variable was TD status and the predictors evaluated were age, gender, and DRD3 gly allele carrier status (Table 5A). There were significant effects of age ($p < .0001$), gender ($p = .005$), and DRD3 gly allele ($p = .04$) on TD status. The net effect of DRD3 allele (i.e., the LR test for DRD3 allele was again significant: $\chi^2=5.02$, $df 1$, $p = .02$). Multinomial logistic regressions on DRD3 ser9gly genotype are shown in Table 5B. There were significant effects of age ($p < .0001$), gender ($p = .04$), and DRD3 genotype ($p = .01$) on TD status. The LR test for DRD3 genotype was significant ($\chi^2 = 7.51$, $df 1$, $p = .002$). We conducted further subsidiary analyses in order to identify the most appropriate genotypic model. A co-dominant (additive) model yielded $\chi^2= 4.46$, $df 1$, $p = .03$. A dominant model was not significant ($\chi^2=1.44$, $df 1$, $p = .2$), whereas a recessive model was the strongest ($\chi^2= 6.3$, $p = .01$).

In addition to classifying TD status in a dichotomous fashion, we performed analyses of the relationship of DRD3 genotype to abnormal involuntary movement as a continuous variable rated by the AIMS. Six groups had AIMS ratings (Bonn, Jerusalem, Newcastle, Nithsdale, Toronto African-Americans, and Toronto Caucasians) and were included in the analyses. In ANCOVA analyses, AIMS total score was the dependent variable; group, gender, and DRD3 allele carrier status were the independent variables, and age was a covariate. There were significant main effects of group ($F=6.01$, $df 5$, $p < .0001$), DRD3 allele ($F=4.2$, $df 1$, $p = .04$), and covariate age ($F=56.0$, $df 1$, $p < .0001$), and a significant interaction between group and DRD3 allele ($F=5.4$, $df 5$, $p < .0001$). The same elements were included in a second ANCOVA, except that DRD3 genotype was substituted for DRD3 allele. In this model, the effect of group was significant ($F=5.8$, $df 5$, $p < .0001$) and of the covariate age ($F=53.8$, $df 1$, $p < .0001$). Whereas the main effect of DRD3 genotype was not significant ($F=2.8$, $df 2$, $p = .06$), the interaction between group and DRD3 genotype was ($F=3.7$, $df 9$, $p < .0001$). Post hoc comparison with a contrast test of adjusted mean AIMS total scores (Figure 2) showed significantly higher scores for the gly-gly versus ser-ser ($p < .0001$) and ser-gly ($p = .006$) genotypes.

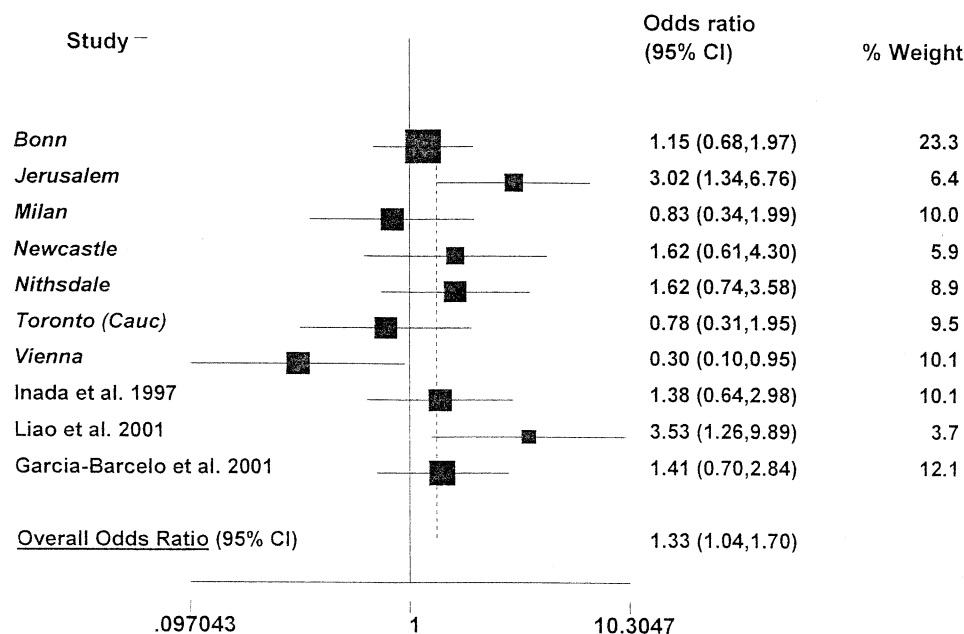


Figure 1. Distribution of odds ratios for TD given DRD3 risk allele (gly) among 7 of the groups included in the current combined analysis and three other published studies. The size of each plot is proportional to the sample size of each group. The Mantel-Haenszel pooled OR is = 1.33 (95% CI 1.04-1.70) and the associated test of OR = is significant ($\chi^2 = 5.12$, df.1, $p = .02$), while the cumulative pooled OR (Lau et al, 1992) is 1.52 with $z = 11.78$ and $p < .001$. The heterogeneity $\chi^2 = 16.97$, df 9, $p = 0.49$.

DISCUSSION

Pharmacogenetics addresses the role of genetic factors in accounting for variability among individuals in response to medication and susceptibility to adverse effects. Case-control association designs are a powerful method for

generating data of this type. Family based designs, though potentially advantageous in reducing spurious effects of population stratification, are difficult to implement in the pharmacogenetic context. In fact, it is highly unlikely that family based samples of sufficient power can be recruited while controlling for both diagnostic

Table 4A. Logistic regression for variables predicting DRD3gly allele carrier status in patients with and without TD from the 8 groups.

Variable	Log Likelihood	Chi Square (df)	Significance	Odds Ratio
Null Model	-1019.8927	—	—	—
Model A	-990.16283	49.43 (8)		
Center				
Bonn				0.92
Jerusalem				1.04
Milan				1.23
Newcastle				1.10
Nithsdale				0.85
Toronto AA				7.96
Toronto Cauc				1.06
Vienna*				—
TD Status				1.26
Model B				
Center	-992.39356	44.97 (7)		
Bonn				0.92
Jerusalem				1.08
Milan				1.24
Newcastle				1.19
Nithsdale				0.89
Toronto AA				8.12
Toronto Cauc				1.07
Vienna*				—
LRtest				
(-2[(LogL model B- (LogL model A)]	—	4.46 (1)	0.034	—

*Dropped due to colinearity; likelihood ratio (LR); tardive dyskinesia (TD).

Table 4B. Multinomial logistic regression for variables predicting DRD3 genotypes in patients with and without TD from the 8 groups

Model (Log-Likelihood)	(Model) Chi Square (df)	Z	p-value	Coefficients Z
Null Model (−769.47)		—	—	—
Model A (−734.14)	70.66 (16)			
[Gen: ser-gly]				
Center				
Bonn		1.429	0.153	0.45
Jerusalem		2.603	0.009	0.92
Milan		2.335	0.020	0.86
Newcastle		1.396	0.163	0.54
Nithsdale		0.350	0.727	0.12
Toronto Cauc		38.514	0.000	22.14
Toronto AA		1.957	0.050	0.73
Vienna*		—	—	—
TD Status		0.451	0.652	.07
[Gen: gly-gly]				
Center				
Bonn		−1.310	0.190	−0.55
Jerusalem		−1.340	0.180	−0.70
Milan		−0.005	0.996	−0.00
Newcastle		0.030	0.976	0.01
Nithsdale		−0.948	0.343	−0.44
Toronto AA**		—	—	−23.10
Toronto Cauc		−0.721	0.471	−0.38
Vienna*		—	—	—
TD Status		2.525	0.012	0.61
Model B (−737.45)	64: 0.3 (14)			
[Gen ser-gly]				
Center				
Bonn		1.437	0.151	.46
Jerusalem		2.645	0.008	.93
Milan		2.343	0.019	.86
Newcastle		1.424	0.155	.55
Nithsdale		0.388	0.698	.14
Toronto AA		—	—	22.0
Toronto Cauc		1.967	0.049	.73
Vienna*		—	—	—
[Gen ser-gly]				
Center				
Bonn		−1.262	0.207	−0.52
Jerusalem		−1.146	0.252	−0.59
Milan		0.050	0.960	−0.02
Newcastle		0.204	0.839	0.10
Nithsdale		−0.706	0.480	−0.33
Toronto AA		40.589	0.000	−23.18
Toronto Cauc		−0.663	0.508	−0.35
Vienna*		—	—	—
LR test (−2[(LogL model B)−(LogL model A)])	—	−6.62 (2)	0.037	—

Likelihood ratio (LR); Tardive dyskinesia (TD).

* Dropped due to colinearity.

** There are no non-TD subjects with gly-gly genotypes in the Toronto African American sample.

and drug exposure elements of the phenotype. Case control designs are not immune to power consideration, as gene effects are likely to be small and interactions among genes will need to be studied, both requiring large samples. Therefore, there is frequently a need to pool samples from different centers and even from different pop-

ulations. Though increasing power, this can lead to spurious results if allele frequencies for the genes being studied vary among the populations. On the other hand, there is a distinct advantage to studying pharmacogenetic issues across populations, so it can be determined to what extent findings are generally applicable.

Table 5. Logistic Regression Analysis for Variables Predicting Tardive Dyskinesia Phenotype in Patients from the 8 Groups**A. Allelic Effect**

Null Model	LogL = .1081.06			LR chi ² =		p
Model A	LogL = .1003.82			LR chi ² (3 df = 96.4)		p < .0001
	Variables	Odds Ratio	Std. Error	[95% Conf. Interval]	Z	p > z
	Age	1.03	.004	1.03–1.04	9.161	0.000
	Sex	1.40	.154	1.13–1.74	3.101	0.002
	DRD3gly Allele	1.28	.142	1.03–1.59	2.243	0.025
Model B	LogL = –1.006.33–			LR chi ² (2 df) = 91.4		p < .001
	Variables	Odds Ratio	Std. Error	[95% Conf. Interval]	Z	p < z
	Age	1.03	.004	1.02–1.04	9.138	0.00
	Sex	1.40	.150	1.13–1.73	3.092	0.02
	Lrtst			LR chi ² = 5.02		p = .02
	(-2[LogL Model B)–(LogL Model A)					

B. Genotypic Effect

Null Model	LogL = –499.41			LR chi ²		p
Model A	LogL = –499.41			LR chi ² (3 df) = 53.1		p < .0001
	Variables	Odds Ratio	Std. Error	[95% Conf. Interval]	Z	p > z
	Age	1.03	.005	1.03–1.05	6.53	0.000
	Sex	1.41	.220	1.04–1.92	2.23	0.025
	DRD3 genotypes	1.06	.022	1.01–1.10	2.73	0.006
Model B	LogL = –503.16			LR chi ² (2 df) = 45.7		p < .0001
	Variables	Odds Ratio	Std. Error	[95% Conf. Interval]	Z	p > z
	Age	1.03	.005	1.02–1.04	6.46	0.000
	Sex	1.40	.217	1.03–1.91	2.18	0.029
	Lrtst			LR ch ² (1 df) = 7.51)		p = .006
	(-2[LogL Model B)–(LogL Model A)]					

Because it is important to be able to conduct large-scale pharmacogenetic studies in a case control context, appropriate approaches for dealing with potential confounders such as population effects need to be developed. Though it is obviously best to use ethnically homogeneous samples, this is not always possible. One method of ruling out population influences on the markers being studied is to type additional markers that are putatively unrelated to the phenotype and to determine the degree of stratification present in the sample (Pritchard and Rosenberg 1999). When samples from different populations are pooled, it is essential to take population effects into account. A further consideration is that demographic factors such as age and gender, as well as other variables, can influence the phenotype and need to be considered in the analyses.

In the current analysis, we applied a statistical approach, stepwise logistic regression, in order to evaluate the relationship between susceptibility to TD and a serine to glycine polymorphism in the dopamine D3 receptor gene, in eight groups of patients recruited from different populations. The approach allowed us to account for potentially confounding effects of the origin of the patients, and also for other variables such as age and gender that can influence the phenotype. Previously, Lerer et al. (2001) applied this approach to an analysis of the 5-HT_{2C} receptor gene in 513 patients with recurrent major depression (MDD), 649 patients with bipolar disorder (BPD), and 901 normal controls from 9 European countries. We were able to demonstrate significant association of a cys23ser polymorphism in the 5-HT_{2C} receptor gene with MDD and

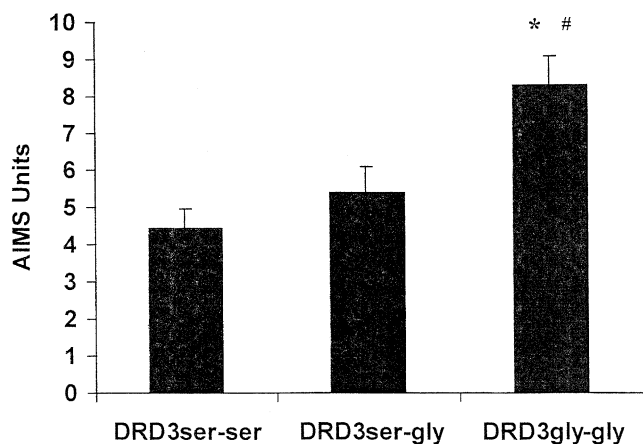


Figure 2. Adjusted mean AIMS scores (bars represent standard error) derived from ANCOVA with age as covariates of patients from 6 groups (Bonn, Jerusalem, Newcastle, Nithsdale, Toronto African Americans, Toronto Caucasians) according to DRD3 ser9gly genotype. *vs. ser-ser, $p < .0001$; #versus ser-gly, $p < .006$.

BPD, in spite of substantial variability in allele frequency among the different populations (considerably greater than that observed for DRD3 ser9 gly in the current study).

In terms of the genes examined, there are two distinct approaches to the design of pharmacogenetic studies. The first employs a classic candidate gene approach and examines genes that have an a priori potential relationship to the phenotype under study. The second approach is atheoretical and involves scanning large numbers of SNPs in order to identify loci associated with the phenotype. Ultimately, whole genome scanning by this method is anticipated (Risch, 2000), although there is considerable debate as to the theoretical underpinnings of the approach (Weiss and Terwilliger 2000) and there are still practical and financial impediments to its application. The present analysis involved data collected on the basis of a candidate gene approach and examined one SNP in a single gene. It demonstrates the small contribution of a single gene to vulnerability to TD and also indicates that this contribution is variable among populations. It is likely, however, that multiple genes contribute to susceptibility to TD, as is anticipated for pharmacogenetic traits in general (Roses 2000).

The results reported here must be considered in the context of significant limitations. From the phenotypic standpoint, evaluations of TD were conducted by a number of different clinicians in different centers. Although inter-rater reliability was achieved within centers, this was not the case between centers. Moreover, different rating scales were used to rate abnormal involuntary movements as a basis for applying the RDC-TD. These limitations are inevitable in a post hoc analysis of data collected in different centers. Nevertheless, there

was standardization in applying the RDC-TD, and five of the eight patient groups were rated with the AIMS. Also, the inclusion of group as a confounding variable in the logistic regression controls for this problem to a certain extent. It should also be noted that genotyping was done in different laboratories. But methods are quite standard, and the margin of error reasonably small.

Overall, the results of this analysis support a contribution of the DRD3 ser9 gly polymorphism to TD susceptibility. The likelihood that patients who carry the gly variant of the gene will manifest TD is increased to a small yet significant degree over those who do not. This conclusion is supported by the results of the meta-analysis, which included the studies that were part of the pooled analysis as well as three additional published studies. It demonstrated a pooled odds ratio of 1.33 and a cumulative odds ratio of 1.52. The conclusion is strengthened by the results of the logistic regression analyses that demonstrated a significant effect of the DRD3 ser9 gly polymorphism over and above the significant contributions of group, age, and gender. Other than preliminary suggestions from in vitro studies regarding an effect of the ser9gly variant on receptor binding (Lundstrom and Turpin 1996), the mechanism whereby DRD3 might be implicated in susceptibility to TD remains unclear and should be studied further in appropriate in vivo models. Besides DRD3, other genes that have been associated with TD include the 5-HT_{2C} receptor gene (HTR2C) (Segman et al. 2000), the 5-HT_{2A} (HTR2A) receptor gene (Segman et al. 2001), cytochrome P450 1A2 (CYP1A2) (Basile et al. 2000), and manganese superoxide dismutase (Hori et al. 2000). These findings have still to be supported by other groups, one replication (Tan et al. 2001) and one non-replication (Basile et al., 2001) having been published for HTR2A, and one non-replication for CYP1A2 (Schulze et al. 2001). It is likely that other associations will be observed and, if replicated, elucidate a polygenic background to TD.

For complex phenotypes such as response to psychotropic drugs and susceptibility to adverse effects (of which TD is an illustrative example), it is likely that several genes will be implicated, each individual gene contributing only a relatively small risk increment. A major challenge for pharmacogenetics is how to conceptualize and best detect the presence or absence of synergistic effects of genes (Phillips 1998; Cordell et al. 2001). These effects are likely to be additive or interactive (epistatic). If the combined effect of two or more genes is greater than the simple sum of the effects of each single gene, then the gene-gene interaction points to an epistatic rather than an additive effect. In a simple formulation, we are interested in seeing whether the observed phenotype is better predicted by a combination of risk genotypes at different loci when they are

jointly present in the same subject rather than when they act independently, as in those subjects where only one of the risk genotypes is present. In this way, the test of a statistical genetic interaction is similar to the concept of epistasis proposed by Fisher (1918). In the context of TD, Segman et al. (2000) demonstrated an additive contribution of *DDRD3* and *HTR2C* to the severity of abnormal involuntary movements (AIMS scores) in schizophrenia patients without evidence for an interaction. Interactions between loci may be within a single gene, as demonstrated for the effect of complex haplotypes of SNPs in the coding region and promoter of the β -2 adrenergic receptor gene on the bronchodilator response to asthma therapy (Drysdale et al. 2000). Interactions may be between loci in different genes as recently observed in a study of genetic susceptibility to breast cancer (Ritchie et al. 2001). In the case of TD, Segman et al. (2002) have demonstrated an interaction between the cytochrome P450 17 α -hydroxylase gene (*CYP-17*) and *DRD3*, an effect of *CYP-17* on AIMS scores being demonstrable only in patients homozygous for the gly allele of *DRD3*.

These considerations indicate that definitive studies on the pharmacogenetics of TD will require large samples. Prospective designs that employ standardized evaluation instruments will reduce heterogeneity among centers in the characterization of the phenotype, enhancing the likelihood of consistent results. As has been demonstrated for *DRD3*, smaller studies can identify positive findings that are subsequently replicated in larger samples. Given the likely small contribution of each gene to susceptibility, however, the potential for type II error is substantial. Certainly, the analysis of interactive effects among several genes will not be possible in small samples, although novel methods of analysis can permit the use of smaller samples than had previously been anticipated (Ritchie et al. 2001). These considerations are generally applicable to the field of pharmacogenetics, and to the pharmacogenetics of psychotropic drugs in particular. In the context of such large, multicenter samples, application of statistical techniques such as those used in the present study, will allow gene effects to be demonstrated over and above the confounding influence of population, demographics, and other confounding variables.

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